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The T2-FLAIR mismatch sign as a predictor of IDH-mutant, 1p/19q-noncodeleted lower-grade gliomas: a systematic review and diagnostic meta-analysis

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Abstract

Objectives To evaluate the diagnostic performance of the T2-FLAIR mismatch sign for prediction of isocitrate dehydrogenase (IDH)-mutant, 1p/19q-noncodeleted lower-grade gliomas (LGGs) and review studies with false positive results.

Methods The MEDLINE and EMBASE databases were searched up to March 13, 2020, to identify articles reporting the diagnostic performance of the T2-FLAIR mismatch sign for prediction of *IDH*-mutant, 1p/19q-noncodeleted LGGs (*IDH*mut-Noncodel) using the search terms (T2 FLAIR mismatch). Pooled sensitivity, specificity, and correlation coefficient for interobserver agreement were calculated.

Results Twelve studies including a total of 1053 patients were included. The median age was 43 (median; range, 14–56). The pooled sensitivity and specificity were 42% (95% CI, 28–58%) and 100% (95% CI, 88–100%), respectively. According to the HSROC curve, the area under the curve was 0.77 (95% CI, 0.73–0.80). Considerable heterogeneity was possible among the studies in terms of both sensitivity and specificity. A threshold effect was suggested and was considered to explain most of the heterogeneity. Four studies reported false positive results for the T2-FLAIR mismatch sign, including dysembryoplastic neuroepithelial tumor, pediatric-type gliomas, and nonneoplastic lesions. The 2 original articles with false positive results showed the highest sensitivities among the 10 studies included in the quantitative analysis, supporting the probability of the threshold effect. The pooled correlation coefficient was 0.87 (95% CI, 0.73–0.94). Conclusions The T2-FLAIR mismatch sign had a high specificity and interobserver agreement for the prediction of *IDH*mut-Noncodel. However, the sign demonstrated low sensitivity, and a few studies with false positive cases were also reported. Key Points

- The pooled sensitivity and specificity of the T2-FLAIR mismatch sign for prediction of IDH-mutant, 1p/19q-noncodeleted lower-grade gliomas were 42% and 100%, respectively.
- Four studies reported false positive results.
- The pooled correlation coefficient was 0.87, suggesting almost perfect interobserver agreement.

Keywords Astrocytoma . Oligodendroglioma . Glioma . Brain neoplasms . Magnetic resonance imaging

Introduction

The revised World Health Organization (WHO) 2016 classification for central nervous system tumors added molecular and genetic features as well as microscopic findings for the classification of diffuse lower-grade gliomas (LGGs) [\[1](#page-9-0)]. According to the classification, based on the mutation status of the isocitrate dehydrogenase (IDH) 1 and 2 genes and the codeletion status of chromosomes 1p and 19q, LGGs are classified into the following: (i) IDH-mutant, 1p/19q-codeleted LGGs (IDHmut-Codel); (ii) IDH-mutant, 1p/19qnoncodeleted LGGs (IDHmut-Noncodel); and (iii) IDH wild-type LGGs (*IDHwt*) [[1\]](#page-9-0). The outcomes of patients with LGGs are known to be stratified across the subtypes, with the worst outcomes associated with IDHwt, the most favorable with *IDH*mut-Codel, and intermediate with *IDH*mut-Noncodel [[2](#page-9-0)].

The T2-FLAIR mismatch sign, initially described by Patel et al, is defined as a complete or near-complete homogeneous high signal intensity on T2-weighted images (T2WI) and a relative suppression of the signal intensities on the fluidattenuated inversion recovery (FLAIR) sequence [[3\]](#page-9-0). Following the initial description by Patel et al [[3](#page-9-0)] and subsequent validation by Broen et al [[4\]](#page-9-0), the T2-FLAIR mismatch sign was reported to demonstrate a near-perfect positive predictive value and specificity for the prediction of IDHmut-Noncodel [\[3](#page-9-0)–[7](#page-9-0)]. However, it also exhibited relatively lower sensitivities (from 22 to 89%) and a few studies even reported false positive results $[3, 8, 9]$ $[3, 8, 9]$ $[3, 8, 9]$ $[3, 8, 9]$ $[3, 8, 9]$ $[3, 8, 9]$. In order to use it as a biomarker for IDHmut-Noncodel, it is important to ascertain that the T2- FLAIR mismatch sign has a high positive predictive value with little or no false positive results and to identify conditions in which false positive results are observed.

Although two reviews were recently published dealing with this subject $[10, 11]$ $[10, 11]$ $[10, 11]$ $[10, 11]$, none of them has quantitatively evaluated the diagnostic performance and interobserver agreement of the T2-FLAIR mismatch sign. Therefore, this systematic review and meta-analysis was performed to evaluate the diagnostic performance of the T2-FLAIR mismatch sign for

the prediction of IDHmut-Noncodel and to review the studies that reported the false positive results.

Materials and methods

This meta-analysis was performed according to the Preferred Reporting Items for Systematic Reviews and Meta-analysis of Diagnostic Test Accuracy Studies (PRISMA-DTA) statement [\[12](#page-9-0)].

Literature search

The MEDLINE and EMBASE databases were searched up to March 13, 2020, to identify articles that reported the diagnostic performance of the T2-FLAIR mismatch sign for the prediction of IDHmut-Noncodel. The search term used was (T2 FLAIR mismatch), and the results were limited to English publications. The references provided in the selected articles were also screened for identification of additional eligible studies.

Study selection

Inclusion criteria

Articles were included based on the fulfillment of all the following criteria: (1) patients with pathologically confirmed LGGs; (2) MRI showing the presence or absence of the T2- FLAIR mismatch sign as an index test; (3) histopathological examination with the IDH mutation and 1p/19q codeletion status as a reference standard; (4) sufficient data for reconstruction of 2×2 tables in terms of the diagnostic performance of the T2-FLAIR mismatch sign.

Conference abstracts with sufficient data for 2×2 tables were included in the meta-analysis. Case reports or case series including 5 patients or fewer were included for the purpose of qualitative but not quantitative synthesis.

Exclusion criteria

Articles were discarded if they fulfilled any of the following criteria: (1) reviews, letters, guidelines, editorials, or errata; (2) insufficient data for the reconstruction of 2×2 tables; (3) studies with overlapping cohorts.

Two authors (S.I.P. and C.H.S. [1 and 7 years of experience in performing systematic reviews and meta-analyses, respectively]) independently evaluated the eligibility of the articles, and any disagreement was resolved via discussion with a third author (H.S.K., 22 years of experience in neurooncology).

Data extraction and quality assessment

A standardized form was used for extracting the following data.

- 1. Study characteristics—authors, year of publication, institution, country of origin, study period, study design (prospective vs. retrospective), and type of enrollment (consecutive vs. non-consecutive).
- 2. Patient and clinical characteristics—number of patients, males:females, mean or median age, age range, and number of patients with IDH-mutant gliomas, IDHwt, IDHmut-Codel, IDHmut-Noncodel.
- 3. Technical characteristics of MRI—magnetic field strength (T), vendor, scanner, head coil, and pulse sequences.
- 4. Interpretation of MRI—number of readers, reader experience, blinding to IDH mutation and 1p/19q codeletion status, definition of the T2-FLAIR mismatch sign, and interobserver agreement.
- 5. Reference standard—type of IDH genes tested (only IDH1 or both IDH1 and IDH2), IDH mutation testing method, and 1p/19q codeletion testing method
- 6. Diagnostic performance of the T2-FLAIR mismatch sign for the prediction of IDHmut-Noncodel—number of true positives, false positives, false negatives, and true negatives.

The quality of the selected studies was evaluated using Quality Assessment of Diagnostic Accuracy Studies-2 (QUADAS-2) [[13\]](#page-9-0). The quality was independently evaluated by the two authors (S.I.P. and C.H.S.), and disagreements were resolved through discussion with the third author (H.S.K.).

Data synthesis and analysis

The diagnostic performance of the T2-FLAIR mismatch sign for prediction of IDHmut-Noncodel was regarded as the primary outcome of this meta-analysis. Two by two tables were constructed for each study for the calculation of pooled estimates.

Pooled sensitivity and specificity were calculated using the bivariate and hierarchical summary receiver operating characteristic (HSROC) models [\[14](#page-9-0)–[16\]](#page-9-0). Coupled forest plots and HSROC curves were constructed for illustrating the results. Deeks' funnel plot with asymmetry test was used for evaluating the publication bias and its statistical significance [\[17](#page-9-0)]. For interobserver agreement, pooled correlation coefficient was calculated using a random-effects model with Fisher's Z transformation of correlations [[18](#page-9-0)].

The presence or absence of heterogeneity was assessed using the following methods: (1) visual evaluation of the difference in area between the 95% confidence and prediction regions of the HSROC curve; (2) Cochran's Q test, with $p < 0.05$ suggestive of

the possibility of heterogeneity; and (3) Higgins' l^2 index, with a value $> 50\%$ suggestive of the possibility of heterogeneity [\[19\]](#page-9-0). The threshold effect was assessed using the following methods: (1) visual evaluation of the coupled forest plot to observe the correlation between sensitivity and false positive rate among the studies; and (2) Spearman correlation coefficient between sensitivity and false positive rate, with a value ≥ 0.6 suggestive of a considerable threshold effect [\[20](#page-9-0)].

For statistical analyses, the "midas" and "metandi" modules in Stata 15.0 (StataCorp LP) and the "meta" and "mada" packages in the R software version 3.6.2. (R Foundation for Statistical Computing) were used. $p < 0.05$ was considered to denote statistical significance.

Results

Literature search

The study selection process is illustrated in Fig. [1](#page-3-0). The initial literature search yielded 83 articles: 28 from the MEDLINE and 55 from the EMBASE databases, respectively. After removing 30 duplicate articles, the remaining 53 articles were screened on the basis of their title and abstract, and 38 articles were excluded. Full texts of the remaining 15 articles were obtained and reviewed, and 3 articles were again excluded because they were either conference abstracts with insufficient data for the reconstruction of 2×2 tables $(n = 2)$ [[21](#page-9-0), [22\]](#page-9-0) or outside the field of interest $(n = 1)$ [\[23\]](#page-9-0). Finally, 12 studies (8) original articles, 3 conference abstracts, 1 case series, and 1 case report) that involved the reports of 1053 patients were included in this study [[3](#page-9-0)–[9](#page-9-0), [24](#page-10-0)–[28\]](#page-10-0).

Characteristics of included studies

The characteristics of the included patients and studies are shown in Tables [1](#page-4-0) and [2,](#page-5-0) respectively. The median number of patients for each study was 106 (range, 1–154). The mean or median age of the patients per study was 43 (median; range, 14–56). The median number of patients with IDHmut-Noncodel was 38 (range, 0–110). Nine studies were retrospective in design [\[3](#page-9-0)–[9,](#page-9-0) [26](#page-10-0), [27\]](#page-10-0), whereas the other 3 did not report the study design [\[24](#page-10-0), [25](#page-10-0), [28\]](#page-10-0). Patient enrollment was consecutive in 3 studies [\[5](#page-9-0), [24,](#page-10-0) [25\]](#page-10-0), non-consecutive in 2 studies [\[26,](#page-10-0) [27\]](#page-10-0), and not reported in the remaining 7 studies [\[3](#page-9-0), [4](#page-9-0), [6](#page-9-0)–[9](#page-9-0), [28\]](#page-10-0).

Among the studies included in the meta-analysis, 9 reported 2 readers [\[3](#page-9-0)–[5,](#page-9-0) [7](#page-9-0)–[9](#page-9-0), [24](#page-10-0), [25,](#page-10-0) [28\]](#page-10-0), whereas 1 did not submit any reader count $[6]$. Among the 9, the reader experience was reported in 4 studies and ranged from 2 to 19 years [[3](#page-9-0)–[5](#page-9-0), [9\]](#page-9-0). The readers were blinded to the reference standard in 6 studies [\[3](#page-9-0)–[5,](#page-9-0) [7](#page-9-0)–[9\]](#page-9-0), whereas no such information was reported in the remaining 4 studies [\[6,](#page-9-0) [24,](#page-10-0) [25,](#page-10-0) [28\]](#page-10-0). Seven studies [\[3](#page-9-0)–[6](#page-9-0), [8,](#page-9-0) [9,](#page-9-0) [24](#page-10-0)] used the same definition of the T2-FLAIR mismatch sign as

initially described by Patel et al as follows: (i) complete or nearcomplete homogeneous hyperintensity on T2WI and (ii) a hypointensity on FLAIR except for a hyperintense rim [\[3](#page-9-0)]. A subjectively determined proportion of T2-FLAIR mismatch of > 50% was considered confirmation of the T2-FLAIR mismatch sign in 1 study [\[7](#page-9-0)]. Two conference abstracts did not provide the definition of the T2-FLAIR mismatch sign [\[25](#page-10-0), [28\]](#page-10-0).

The magnetic field strength of the MRI machine was reported in 4 studies [[4,](#page-9-0) [6](#page-9-0), [7,](#page-9-0) [9](#page-9-0)]: 3 T in 2 studies [[6](#page-9-0), [9](#page-9-0)] and 1.5 T or 3 T in 2 studies [\[4](#page-9-0), [7\]](#page-9-0). Details of MRI machines and pulse sequences are shown in Table [2](#page-5-0).

Quality assessment

A quality assessment summary of the included studies using the QUADAS-2 tool is shown in Supplemental Fig. 1. The overall quality of the included studies was moderate.

With regard to patient selection, 8 studies indicated an unclear risk of bias as they failed to mention the method of patient enrollment (consecutive or not) [\[3](#page-9-0), [4](#page-9-0), [6](#page-9-0)–[9](#page-9-0), [26](#page-10-0), [28](#page-10-0)]. One study caused high concern regarding applicability as the authors had included patients with dysembryoplastic neuroepithelial tumor (DNET) as well as with LGGs [\[28](#page-10-0)].

With regard to the index test, 6 studies were considered to have an unclear risk of bias, as they did not mention whether the readers had been blinded to the reference standard [\[6,](#page-9-0) [24](#page-10-0)–[28\]](#page-10-0). One study was considered as being unclear regarding applicability, as the authors had used a different criterion for the T2-FLAIR mismatch sign (> 50% area of T2-FLAIR mismatch) [\[7](#page-9-0)].

With regard to the reference standard, all the studies presented an unclear risk of bias as they failed to mention whether they had performed a blinded review of the molecular classification. One study was considered to have high concerns regarding applicability, as it used a different definition of the target condition (i.e., IDH mutation with concomitant TP53/ ATRX inactivation, or the absence of 1p/19q codeletions and/ or *TERT* p mutations) [\[8](#page-9-0)].

With regard to the flow and timing, all the studies were considered to have an unclear risk of bias because they did not mention the imaging to surgery intervals.

Diagnostic performance of the T2-FLAIR mismatch sign

Diagnostic performance of the T2-FLAIR mismatch sign

For evaluating the diagnostic performance of the T2-FLAIR mismatch sign for the prediction of IDHmut-Noncodel, 10 studies with 11 cohorts were evaluated [\[3](#page-9-0)–[9](#page-9-0), [24,](#page-10-0) [25](#page-10-0), [28](#page-10-0)]. The coupled forest plot is presented in Fig. [2.](#page-6-0) The sensitivity, specificity, and diagnostic accuracy of the individual studies ranged from 10.9 to 89.5%, 69.2 to 100%, and 13.3 to 88.9%, respectively. The pooled sensitivity and specificity were 42% (95% CI, 28–58%) and 100% (95% CI, 88–100%), respectively. According to the HSROC curve, the area under the curve was 0.77 (95% CI, 0.73–0.80) (Fig. [3](#page-7-0)).

Considerable heterogeneity was possible among the studies in terms of both sensitivity and specificity according to the

Author (year of publication)	Institution		Period	Study design	Consecutive enrollment	No. of patients $\binom{n}{k}$	Mean age (years)	range Age	Male:female
Kinoshita M, et al (2020) [6] Batchala PP, et al (2019) [5] Juradi TA, et al (2019) [8]	University of Virginia Health System, USA Osaka International Cancer Institute, Japan	Massachusetts General Hospital, USA; University Hospital	2010-2017 \lesssim \mathbb{E}	Retrospective Retrospective Retrospective	Y es \tilde{A} \mathfrak{Z}	106 133	38.5^a \tilde{A} $\frac{4}{5}$	$17 - 70$ $20 - 81$ $\tilde{\mathbf{z}}$	53:53 $\tilde{\ge}$ 4.5
Broen MPG, et al (2018) [4] Lasocki A, et al (2018) [7] Lee MK, et al (2019) [9]	Peter MacCallum Cancer Centre, Australia Asan Medical Center, South Korea Dresden, Germany	Erasmus MC, University Medical Center Rotterdam and Maastricht University Medical Center, Netherlands	Three different cohorts $(2003 -$ 2010 August-2016 August 2015 May-2017 May $2013 - 2015 -$	Retrospective Retrospective Retrospective	\lesssim \tilde{A} \lessapprox	110 154 59	$\stackrel{\triangle}{\simeq}$ \$6	$19 - 82$ $20 - 82$ $\stackrel{\triangle}{\simeq}$	56:54 86:68 $\stackrel{\triangle}{\approx}$
Patel SH, et al (2017) (1) [3] Patel SH, et al (2017) (2) [3] Conference abstracts	NYU Langone Medical Center, USA TCGA/TCIA database		2011-2014 Ź	Retrospective Retrospective	\tilde{z} \lessapprox	125 \otimes	45.5^a \lesssim	$20 - 75$ \tilde{M}	62:63
Foltyn M, et al (2019) [24] Galldiks N, et al (2019) [25] Onishi S, et al (2019) [28]	Japan	Chugoku Cancer Center, Hiroshima University Hospital, University of Cologne, Research Center Juelich, Germany National Hospital Organization Kure Medical Center and University of Heidelberg Medical Center, Germany	\lesssim ≸ \lessapprox	\mathbb{E} $\stackrel{\triangle}{\geq}$ \lesssim	Yes Yes \lessapprox	113 134 $\overline{4}$	\lessapprox \lesssim \tilde{A}	\lessapprox $\tilde{\mathbf{z}}$ \tilde{z}	≨ ≨ ≨
Studies not included in meta-analysis Niemeyer B, et al (2018) [27] Johnson DR, et al (2019) [26]	Instituto Estadual do Cérebro Paulo Multi-institutional case series	Niemeyer, Brazil	\lesssim \lessapprox	Retrospective Retrospective	$\rm \simeq$ \tilde{z}	\sim $-$	$14^{\rm a}$ 33	$2 - 44$ \lessapprox	5:0 \overline{c}
IDH-Mutant glioma (n)	IDHmut-Codel (n)	IDHmut-Noncodel (n)	ЮH DH wt (n)		Reference test (IDH mutation)		Reference test (1p/19q-codeletion)		
106 142 124 65	46 56 42 67	$\begin{smallmatrix} 50 \\ 82 \\ 0 \end{smallmatrix}$ 75	DH1/2 IDH1/2 IDH1/2 IDHI IDHI 45 C 500	≨	NGS, Sanger sequencing, or IHC Standard genomic sequencing IHC, DNA pyrosequencing NGS or Sanger sequencing		FISH FISH FISH FISH ≸		
Conference abstracts 102 59 53	34 $\overline{31}$ $\overline{21}$	38 68	IDH1/2 DHI IDHI \circ \sim 33	IHC, NGS ЭH	Whole-exome sequencing		PCR loss of heterozygosity Affymetrix SNP6.0 arrays FISH		
\lessapprox Z 65	$\stackrel{\prec}{\geq}$ $_\odot$ \lessapprox	$\overline{110}$ 65 $\overline{18}$	\lessapprox $Z \nleq$ Σ \otimes Σ	\tilde{M} $\tilde{\mathbf{z}}$			≨≨≨		
Studies not included in meta-analysis $1b$		0	$Z \nleq$ $\frac{1}{2}$ 0	≨≨			$Z \nleq$		
type lower-grade glioma; IHC, immunohistochemistry; NGS, next- NA, not available; IDH, isocitrate dehydrogenase; IDHmut-Codel, ^b Tested in 2 patients a Median age			IDH-mutant, 1p/19q-codeleted lower-grade glioma; IDHmut-Noncodel, IDH-mutant, 1p/19q-noncodeleted glioma; IDHw, IDH wild- generation sequencing; FISH, fluorescence in situ hybridization; PCR, polymerase chain reaction						

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 $^{\rm d}$ Low-grade isomorphic astrocytom
a Low-grade isomorphic astrocytoma

⁶ IDH mutation with concomitant TP53/ATRX inactivation, or the absence of 1p/19q codeletions and/or TERTp mutations IDH mutation with concomitant TP53/ATRX inactivation, or the absence of 1p/19q codeletions and/or TERTp mutations

weighted imaging; DSC, dynamic susceptibility contrast imaging

Fig. 2 Coupled forest plot of pooled sensitivity and specificity for evaluating the diagnostic performance of the T2-FLAIR mismatch sign for the prediction of IDHmut-Noncodel

Cochran's Q test and Higgins' l^2 statistics (p < 0.01 for Q test, $I^2 = 91.8\%$ for sensitivity; $p < 0.01$ for Q test, $I^2 =$ 89.7% for specificity). A large difference in the area between the 95% confidence and prediction regions on the HSROC curve also indicated possible heterogeneity among the studies (Fig. [3](#page-7-0)).

The near V-shape of the coupled forest plot on visual evaluation and the Spearman correlation coefficient (between sensitivity and false positive rate) of 0.63 (95% CI, 0.05–0.89) suggested the possibility of a threshold effect, which accounted for most of the heterogeneity among the studies. The Deeks funnel plot with an asymmetry test revealed a low probability of publication bias $(p = 0.11)$ (Supplemental Fig. 2).

Studies with false positive results

The T2-FLAIR mismatch sign provided false positive results for the prediction of IDHmut-Noncodel in 4 studies (2 original articles $[8, 9]$ $[8, 9]$ $[8, 9]$ $[8, 9]$, 1 conference abstract $[28]$ $[28]$, and 1 case series $[26]$ $[26]$; 3 of them were included in the meta-analysis $[8, 9, 28]$ $[8, 9, 28]$ $[8, 9, 28]$ $[8, 9, 28]$ $[8, 9, 28]$.

Among the 10 studies included in the quantitative analysis, 2 original articles (Lee et al [[9\]](#page-9-0) and Juratli et al [[8\]](#page-9-0)) demonstrated the highest sensitivities of 89.5% and 73.2%, respectively (Fig. 2). Lee et al reported that 39.1% (18/46) of IDHmut-Codel and 15.6% (7/45) of IDHwt cases exhibited the T2-FLAIR mismatch sign [\[9](#page-9-0)]. In the study by Juratli et al, 28.6% (12/42) of IDHmut-Codel cases were positive for the T2-FLAIR mismatch sign, whereas no IDHwt case was positive for the T2- FLAIR mismatch sign [[8\]](#page-9-0).

The study by Onishi et al included patients with both DNET and LGGs [\[28](#page-10-0)], and 8 out of 11 patients with DNET were positive for the T2-FLAIR mismatch sign. No LGG except IDHmut-Noncodel exhibited the T2-FLAIR mismatch sign. In the case series by Johnson et al, there was 1 adult patient with IDHmut-Codel, while the other 4 patients were children or young adults with pediatric-type gliomas or nonneoplastic lesions [\[26](#page-10-0)].

Fig. 3 Hierarchical summary receiver operating characteristic (HSROC) curve of diagnostic performance of the T2-FLAIR mismatch sign for the prediction of IDHmut-Noncodel

Interobserver agreement

For testing the interobserver agreement, 9 studies with 10 cohorts were evaluated using a random-effects model for the calculation of pooled correlation coefficient [\[3](#page-9-0)–[5,](#page-9-0) [7](#page-9-0)–[9,](#page-9-0) [24](#page-10-0), [25,](#page-10-0) [28\]](#page-10-0). The interobserver agreement among the individual studies ranged from 0.56 to 1, and the pooled correlation coefficient was 0.87 (95% CI, 0.73–0.94), suggesting almost perfect agreement (Fig. 4). Cochran's Q test ($p < 0.01$) and the Higgins I^2 statistic indicated the possibility of heterogeneity.

There was no publication bias $(p = 0.11)$ (Supplemental Fig. 3).

Discussion

This systematic review and meta-analysis evaluated the diagnostic performance of the T2-FLAIR mismatch sign for prediction of IDHmut-Noncodel. The pooled sensitivity and specificity were 42% (95% CI, 28–58%) and 100% (95% CI, 88–100%), respectively. False positive cases of the T2- FLAIR mismatch sign were observed in only 4 studies [[8](#page-9-0), [9,](#page-9-0) [26,](#page-10-0) [28](#page-10-0)]. The pooled correlation coefficient was 0.87 (95% CI, 0.73–0.94). Therefore, the T2-FLAIR mismatch sign for prediction of IDHmut-Noncodel exhibited high specificity and interobserver agreement. However, it also demonstrated low sensitivity, and a few studies showed false positive cases.

As reported in multiple studies and this meta-analysis, the T2-FLAIR mismatch sign has the advantage of near-perfect specificity for predicting *IDH* mut-Noncodel. Thus, it can serve as a radiogenomic biomarker in the diagnosis of LGGs. As the IDHmut-Noncodel group has intermediate outcomes among LGGs [\[3](#page-9-0)], pretreatment identification of the T2- FLAIR mismatch sign may aid in deciding management options [[4,](#page-9-0) [29\]](#page-10-0). Furthermore, as IDH-mutant astrocytomas with even small tumor remnants after surgery were shown to have worse outcomes, patients displaying the T2-FLAIR mismatch sign in preoperative MRI may need to be treated more radically (i.e., towards gross total resection) [\[4](#page-9-0), [30](#page-10-0)–[32](#page-10-0)]. However, the low sensitivity of the T2-FLAIR mismatch sign should be considered when using it for imaging classification of a brain tumor—its absence does not exclude the possibility of IDHmut-Noncodel. Despite its low sensitivity, the high specificity (i.e., positive predictive value) of the T2-FLAIR mismatch sign makes it a useful predictor for IDHmut-Noncodel.

Apart from its high specificity, the T2-FLAIR mismatch sign presents another advantage: that it does not require any advanced imaging techniques such as perfusion-weighted MRI, including DCE (dynamic contrast-enhanced imaging) or DSC (dynamic susceptibility contrast imaging), or MR spectroscopy. The wide availability and high interobserver agreement of the T2-FLAIR mismatch sign make it a useful biomarker for the preoperative diagnosis and classification of LGGs.

Regarding the false positive results, Jain et al argued that they may have been caused by the non-strict application of the T2-FLAIR mismatch sign [[11](#page-9-0)]. It is supported by the results of this meta-analysis that there was the threshold effect accounting for the heterogeneity among the studies and that the 2 original articles with false positive results demonstrated the highest sensitivities among the included studies [\[8](#page-9-0), [9\]](#page-9-0). As Patel et al initially described, the T2-FLAIR mismatch sign requires (i) complete or near-complete homogeneous

hyperintensity on T2WI and (ii) a hypointensity on FLAIR except for a hyperintense rim [[3](#page-9-0)]. Thus, the homogeneity of signal intensity on T2WI and a hyperintense rim on FLAIR are also required for the presence of the T2-FLAIR mismatch sign. The presence of only a discrepancy in the signal intensity on T2WI and FLAIR (such as in a cyst) in itself is insufficient to confirm the T2-FLAIR mismatch sign.

The exact mechanism underlying the presence of the T2- FLAIR mismatch sign in IDHmut-Noncodel cases is still to be established. Patel et al reported that higher prevalence of abundant microcysts was observed in tumors with a positive T2- FLAIR mismatch sign, but it failed to reach statistical significance ($p = 0.128$) [\[3](#page-9-0)]. One possible explanation is that the T2-FLAIR mismatch sign may reflect the cellularity of the tumor. Further studies with MRI-pathology correlation may provide a clue on the exact mechanism.

The results of this meta-analysis should be applied with caution to routine image interpretation. In the majority of included studies, the target subjects were limited to those with proven LGGs, and not all space-occupying lesions undergoing tissue confirmation. Thus, the observed high specificity was not proven against tumors except LGGs or nontumorous lesions. For example, patients with dysembryoplastic neuroepithelial tumors (DNETs) were also reported to show the T2-FLAIR mismatch sign [\[33\]](#page-10-0). In that study, most patients with a positive T2-FLAIR mismatch sign were aged < 20 years, similarly to the study by Johnson et al [\[26](#page-10-0), [33](#page-10-0)]. Therefore, the T2-FLAIR mismatch sign should be applied in patients only after considering their age and the possibility of LGGs.

Two questions remained unclear in this systematic review and meta-analysis. First, regarding the outcomes of patients depending on the presence or absence of the T2-FLAIR mismatch sign, only 2 studies reported no significant differences between the mismatch positive and mismatch negative groups [\[3](#page-9-0), [8](#page-9-0)]. Second, regarding the diagnostic performance of the T2-FLAIR mismatch sign for the prediction of IDHmut-Noncodel according to the histologic grade of LGGs, 2 studies provided sensitivities of the T2-FLAIR mismatch sign separately for grade 2 and grade 3 groups: Broen et al (48.6% (34/ 70) for grade 2, 80% (4/5) for grade 3) [[4](#page-9-0)] and Juratli et al (68.7% (22/32) for grade 2, 76% (38/50) for grade 3) [[8\]](#page-9-0). Future studies with more patients might give clues to these questions.

This study has the following limitations. First, the number of analyzed studies was small, with 3 of them being conference abstracts [\[24,](#page-10-0) [25](#page-10-0), [28\]](#page-10-0); thus, additional analyses could not be performed. However, the high specificity of the T2-FLAIR mismatch sign was consistently demonstrated among these studies except for a few exceptions. Second, heterogeneity was observed among the studies and the diagnostic performance profiles of individual studies were highly variable. The heterogeneity was partly attributed to the threshold effect

(inverse relationship between sensitivity and specificity among studies) but could not be analyzed further. Third, the selection criteria among the studies had slight differences (e.g., inclusion of IDHwt, inclusion of contrast-enhanced tumors).

Conclusions

The T2-FLAIR mismatch sign had high specificity and interobserver agreement for the prediction of IDHmut-Noncodel. However, low sensitivity was observed, and a few studies had false positive cases.

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Methodology

• systematic review and meta-analysis

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